A change detection measure based on a likelihood ratio and statistical properties of SAR intensity images

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With its weather- and illumination-independent characteristics, synthetic aperture radar (SAR) has become an important tool for change detection. There are two critical steps in SAR image change detection: designing a change detector and choosing a decision rule. Given a measure from a change detector, the change detection results could be sensitive to the decision rule, such as the selection of a threshold. This letter presents a change detection measure based on a likelihood ratio and the statistical distribution of SAR intensity images. The likelihood ratio is defined as the ratio between the joint probability density functions (PDFs) of a pair of SAR images. Under the condition that both PDFs follow the gamma distribution, the histogram of this change detection measure deduced from the likelihood ratio has a single and steep peak that can be used to reliably and easily determine the change detection threshold. Analyses of SAR image pairs from different platforms show that the proposed change detection measure is simple and effective in detecting changes.

1. Introduction

Change detection is a process of acquiring time-dependent information of objects of interest from the same scene observed at different times. Synthetic aperture radar (SAR) has been an important technique for change detection for various applications such as environmental monitoring (Wang and Allen 2008, Makkasorn et al. 2009), agricultural surveys, urban studies, forest monitoring and disaster evaluation (Bovolo and Bruzzone 2007). Changes in a SAR scene are manifested mainly as differences of intensity or local texture among multi-temporal SAR images. Image differencing is one of the most widely used change detection algorithms for a variety of geographical environments (Singh 1989). Image ratioing is another change detection method which is also as simple and quick as image differencing. It is widely used in SAR image change detection because of multiplicative noise reduction when data are ratioed on a pixel-by-pixel basis. It is well known that, intrinsic to all active imaging systems, SAR images suffer from speckle noise due to coherence from distributed targets (Franceschetti and Lanari 1999). The variety of the noise and imaging conditions

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in the data acquisition process affects the performance and stability of SAR change
detection. In particular, a simple threshold of 0 for image differencing and 1 for image
ratioing may not be appropriate to identify change from no-change. In this letter we
propose a change detection measure that is based on a likelihood ratio of SAR inten-
sity images and a theoretical analysis of SAR statistical properties. With the proposed
change detection measure, one can estimate an optimum threshold that differentiates
the changes from the non-changed areas.

2. The proposed change detection measure

There are two key steps in classical SAR image change detection (Inglada and Mercier
2007). First is the development of a change detection measure, and the second is to
make a decision with this measure. Image differencing appears to be an appropri-
ate measure in change detection if the data are affected by additive Gaussian noise
(Coppin and Bauer 1996). However, SAR speckle is a multiplicative noise, and thus
image ratioint of the multi-temporal radar intensity data is better adapted to the
statistical features of SAR data than image differencing (Rignot and Zyl 1993). An
alternative change detection measure widely used in SAR change detection is the log-
ratio (Bazi et al. 2006), which converts the linear scale of SAR data to a logarithmic
scale before the differencing operation, and thus transforms the multiplicative noise
into additive noise.

The results after image differencing or ratioing are called ‘residual data’ (Jha and
Unni 1994). The choice of a threshold differentiating the change from no-change in
unsupervised change detection can be made automatically on the basis of residual
data. The automatic selection of thresholds is of particular interest in SAR change
detection (Bruzzone and Prieto 2000, Rogerson 2002, Moser and Serpico 2006). Based
on the likelihood ratio and the gamma distribution of SAR intensity images, a sta-
tistical hypothesis testing is proposed for change detection of SAR data (Conradsen
et al. 2003). With this hypothesis testing, a new change detection measure is deduced,
which is not only as simple as image differencing, but also reliable and provides an
easy method for selecting a threshold.

In this letter, each pixel in a SAR image is assumed to be independent of the other
pixels. If \( x \) expresses the intensity value of a pixel in a homogenous region of an \( L \)-look
SAR image, the probability density function (PDF) of \( x \) is (Dekker 1998, Oliver and
Quegan 2004)

\[
P(x) = \frac{1}{\Gamma(L)} \left( \frac{L}{u} \right)^L x^{L-1} \exp \left( -\frac{Lx}{u} \right)
\]

where \( \Gamma(\cdot) \) is a gamma function. From the moment estimation, the mean \( (E) \) and
variance \( (D) \) of \( x \) are as follows:

\[
E(x) = u, \quad D(x) = \frac{u^2}{L}
\]

It can be assumed that there are \( M \) groups of SAR images, and each group contains
a SAR image after an \( L \)-look processing. Given a pixel \( x_{ij} \) in a homogenous area,
which is surrounded by \( N-1 \) adjacent pixels that follow the same PDF, the joint PDF
of \( x_{ij} (i = 1, 2, \ldots, N, \ j = 1, 2, \ldots, M) \) can be expressed as
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\[ P(X) = P(x_{11}, x_{12}, \ldots, x_{NM}) = \prod_{i=1}^{N} \prod_{j=1}^{M} \frac{1}{\Gamma(L)} \left( \frac{L}{u_j} \right)^L x_{ij}^{L-1} \exp \left( -\frac{x_{ij}L}{u_j} \right) \]  

(3)

where \( u_j \) is the average intensity of pixels in the homogenous area of each image.

When \( M = 2 \), the first image is regarded as the so-called master image, and the second the slave for change detection. Using the moment estimation in equation (2), the change detection problem is addressed by using the following binary hypothesis test (Chatelain and Tourneret 2007):

\[ H_0 \) (unchanged pattern): \( u_1 = E(x_1) = u_2 = E(x_2) = u_0 \]

\[ H_1 \) (changed pattern): \( u_1 = E(x_1) \neq u_2 = E(x_2) \]

where \( x_1 \) and \( x_2 \) are the pixels in the master and slave images, respectively. Combined with equation (3), the conditional probabilities in a likelihood ratio test can be described as follows:

\[ \lambda_1 = P(X|H_0) = \prod_{i=1}^{N} \frac{1}{\Gamma^2(L)} \left( \frac{L}{u_0} \right)^{2L} x_{i1}^{L-1} x_{i2}^{L-1} \exp \left( -\frac{L}{u_0} (x_{i1} + x_{i2}) \right) \]  

(4)

\[ \lambda_2 = P(X|H_1) = \prod_{i=1}^{N} \frac{1}{\Gamma^2(L)} \left( \frac{L^2}{u_1 u_2} \right)^L x_{i1}^{L-1} x_{i2}^{L-1} \exp \left( -L \left( \frac{x_{i1}}{u_1} + \frac{x_{i2}}{u_2} \right) \right) \]  

(5)

The likelihood ratio, \( \lambda \), is obtained by combining equations (4) and (5):

\[ \lambda = \frac{P(X|H_1)}{P(X|H_0)} = \left( \frac{u_0^2}{u_1 u_2} \right)^{LN} \exp \left( -L \left( \frac{\sum_{i=1}^{N} x_{i1}}{u_1} + \frac{\sum_{i=1}^{N} x_{i2}}{u_2} - \frac{\sum_{i=1}^{N} x_{i1}}{u_0} - \frac{\sum_{i=1}^{N} x_{i2}}{u_0} \right) \right) \]  

(6)

Let \( \eta_1 = \frac{1}{N} \sum_{i=1}^{N} x_{i1}, \eta_2 = \frac{1}{N} \sum_{i=1}^{N} x_{i2} \), then the ratio can be expressed as

\[ \lambda = \frac{P(X|H_1)}{P(X|H_0)} = \left( \frac{u_0^2}{u_1 u_2} \right)^{LN} \exp \left( -NL \left( \frac{\eta_1}{u_1} + \frac{\eta_2}{u_2} - \frac{\eta_1}{u_0} - \frac{\eta_2}{u_0} \right) \right) \]  

(7)

From a maximum likelihood estimation, it is easy to get the estimations of \( u_0, u_1 \) and \( u_2 \) with \( \hat{u}_0 = \arg \max (\lambda_1) \) and \( (\hat{u}_1, \hat{u}_2) = \arg \max (\lambda_2) \):

\[ \left\{ \begin{array}{c}
\hat{u}_0 = \frac{1}{2N} \sum_{i=1}^{N} (x_{i1} + x_{i2}) = \frac{1}{2} (\eta_1 + \eta_2) \\
(\hat{u}_1, \hat{u}_2) = \left( \frac{1}{N} \sum_{i=1}^{N} x_{i1}, \frac{1}{N} \sum_{i=1}^{N} x_{i2} \right) = (\eta_1, \eta_2)
\end{array} \right. \]  

(8)
Thus, equation (7) can be expressed as follows:

$$\lambda = \frac{P(x|H_1)}{P(x|H_0)} = \left( \frac{\eta_1 + \eta_2}{4 \eta_1 \eta_2} \right)^N \exp(0) \begin{cases} > \tau_0 & H_1 \\ < \tau_0 & H_0 \end{cases}$$ (9)

Figure 1 shows the scheme of change detection based on SAR statistical properties. Simplification of equation (9) results in

$$\eta = \eta_1 + \eta_2 - 4 \eta_1 \eta_2 \exp(0) - 2 = \tau$$ (10)

The expression $\eta = (\eta_1 / \eta_2) + (\eta_2 / \eta_1)$ is the new change detection measure, which is a simplified form of the likelihood ratio of the SAR intensity images. The symbol $\tau$ is the new threshold for change detection and $\eta_k$, $k = 1, 2$, are the average values in the sub-homogenous areas of the master and slave SAR images at the same location, respectively. To minimize computational requirements, small windows are used to avoid the necessity of searching for homogenous areas. It is noted that for 8-bit data, when $\eta_1$ equals $\eta_2$, $\eta$ achieves its minimum value $\eta_{\min} = 2$. Therefore, we map $\eta$ to $\eta'$ with the range between 0 and 255. If there is no obvious change identified, then $\eta'$ will be close to 0. Then a threshold is automatically chosen in this range based on the histogram of $\eta'$.

3. Results and analysis

A pair of European remote sensing satellite (ERS-2) SAR images and a pair of airborne SAR images are used to test the proposed change detection measure.

3.1 ERS-2 SAR images

Figure 2(a) and 2(b) shows two co-registered C-band SAR images after radiometric calibration acquired by ERS-2 over the city of Bern, Switzerland, on 20 April and 25 May 1999. The pixel size of the images is about 12.5 m. Flooding over the farmland, evident in figure 2(b), is mapped as the reference image with a visual interpretation in figure 2(c).

For the Bern data set, residual data were calculated using a $3 \times 3$ window, which was assumed to represent a homogenous area. The normalized histogram and its local amplification of the residual data are shown in figure 3(a) and 3(b). The single, sharp peak of the histogram is assumed to represent the unchanged part of the image, and the long tail of the histogram the changed part of the image. Since the unchanged pixels are very close to 0 and concentrated in the narrow peak, the value of the transitional
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Figure 2. A pair of ERS-2 SAR images of Bern, Switzerland: (a) before and (b) after floods. (c) Reference image of change (represented by dark pixels).

Figure 3. (a) Normalized histogram and (b) the local amplification of residual data from the Bern data set. (c) The ratio curve of the histogram and the first point with the ratio under 1.0 are taken as the threshold for change detection.

The point between the peak and the tail can be chosen as the threshold when the histogram curve changes from steep to smooth. The ratios of the histogram at two adjacent grey-level values on the right side of the peak can be used to find the appropriate threshold. The first point with a ratio less than 1.0 is taken as the transitional point. After that point, the histogram enters into an oscillating state and the ratios fluctuate randomly around 1.0. Figure 3(c) is the ratio curve of the histogram. The first point with the ratio less than 1.0 is marked with a circle and its value is 15.

Figure 4(a) is the residual data before thresholding and (b) is the result of the change detection with the selected threshold, which corresponds well with the reference image.

Since log-ratio is a widely used measure in SAR change detection (Bazi et al. 2005, Bovolo and Bruzzone 2005), this letter compares the change detection results with a supervised manual trial-and-error procedure (MTEP) based on a log-ratio detector, which consists of finding the threshold value that minimizes the overall errors, and an unsupervised change detection with Kittler–Illingworth (KI) threshold selection based on general Gaussian (GG) model and the log-ratio detector (Bazi et al. 2005). Table 1 lists the change detection results, the overall error and the kappa coefficient to evaluate these methods. The kappa coefficients and overall errors show that the proposed method is better than that of the GG method and is very close to the result of MTEP.
Figure 4. (a) The residual data with the proposed change detection measure applied to the data shown in figure 2. (b) The change detection result after thresholding high values as changes and mapping them to black pixels.

Table 1. Results achieved by the proposed methods, MTEP and GG method (Bern SAR data set).

<table>
<thead>
<tr>
<th>Method</th>
<th>Correctly identified changed pixels</th>
<th>Correctly identified unchanged pixels</th>
<th>Unchanged pixels classified as changed</th>
<th>Changed pixels classified as unchanged</th>
<th>Overall error</th>
<th>Kappa coefficient</th>
</tr>
</thead>
<tbody>
<tr>
<td>MTEP</td>
<td>1210</td>
<td>127,259</td>
<td>140</td>
<td>272</td>
<td>412</td>
<td>0.853</td>
</tr>
<tr>
<td>GG method</td>
<td>1349</td>
<td>126,950</td>
<td>449</td>
<td>133</td>
<td>582</td>
<td>0.820</td>
</tr>
<tr>
<td>Proposed method</td>
<td>1165</td>
<td>127,288</td>
<td>111</td>
<td>317</td>
<td>428</td>
<td>0.843</td>
</tr>
</tbody>
</table>

3.2 High-resolution airborne SAR images

Figure 5(a) and 5(b) shows two X-band airborne SAR intensity images acquired over a field in Beijing, China, on 4 and 6 April 2004. The pixel size of the images is 0.5 m. The local incident angle was approximately 35°, and the flight azimuth approximately 75°. The numbers and positions of the vehicles on the field were different during the two data acquisition dates. These data sets were only approximately co-registered, with a registration error of about 1–2 pixels. Figure 5(c) is the ground change reference image with a visual interpretation.

Since notable speckle noise is present in these two images, a Frost filter with a $3 \times 3$ window was applied four times before the change detection operation, which resulted in a good trade-off between noise reduction and degradation of the spatial details of the data. Then the same $3 \times 3$ windows were used to replace the homogenous areas to acquire the residual data with the proposed change detection measure. Figure 6(a) and 6(b) shows the histogram and its local amplification of the residual data, and (c) is the ratio curve of the histogram. The grey-level value 18 was selected as the threshold value with the same method in the former experiment.

Figure 7(a) is the residual data before thresholding and (b) is the result of the change detection with the selected threshold, which also corresponds well with the ground reference image.
Figure 5. (a) and (b) A pair of high-resolution airborne SAR images with different vehicle distributions acquired 2 days apart. (c) Reference image of change (represented by dark pixels).

Figure 6. (a) Normalized histogram and (b) the local amplification of residual data from the airborne SAR data set. (c) The ratio curve of the histogram and the first point with the ratio under 1.0 are taken as the threshold for change detection.

Figure 7. (a) The residual data with the proposed change detection measure applied to the data shown in figure 5. (b) The change detection result after thresholding high values as changes and mapping them to black pixels.
Table 2. Results achieved by the proposed methods, MTEP and GG method (airborne SAR data set).

<table>
<thead>
<tr>
<th>Method</th>
<th>Correctly identified changed pixels</th>
<th>Correctly identified unchanged pixels</th>
<th>Unchanged pixels classified as changed</th>
<th>Changed pixels classified as unchanged</th>
<th>Overall error</th>
<th>Kappa coefficient</th>
</tr>
</thead>
<tbody>
<tr>
<td>MTEP</td>
<td>3263</td>
<td>804,911</td>
<td>550</td>
<td>1276</td>
<td>1826</td>
<td>0.780</td>
</tr>
<tr>
<td>GG method</td>
<td>3013</td>
<td>804,989</td>
<td>472</td>
<td>1526</td>
<td>2056</td>
<td>0.750</td>
</tr>
<tr>
<td>Proposed method</td>
<td>3795</td>
<td>803,899</td>
<td>1562</td>
<td>744</td>
<td>2306</td>
<td>0.766</td>
</tr>
</tbody>
</table>

Table 2 lists the evaluation of the change detection result. The kappa coefficients show that the proposed method is more accurate than the GG method and is very close to the result of MTEP. Though the overall error of the proposed method is higher than the benchmark methods, the number of correctly identified changed pixels is much larger than that of the other methods. Most of the false alarm pixels due to misregistration potentially can be removed with a post change-classification procedure.

4. Conclusion

A change detection measure based on the likelihood ratio and statistical properties of SAR intensity images was developed in this letter. With the proposed change detection measure, the histogram of the residual image shows a single, sharp peak for the unchanged areas, which makes it straightforward to select a threshold to differentiate the changed pixels from the unchanged ones. The analyses of two SAR image pairs from different sensors with different resolutions show that the proposed measure can detect the change information from multi-temporal SAR images very well. The change detection results are very close to that from the MTEP and better than that from the unsupervised change detection method with KI thresholding based on the GG model.

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